

RESEARCH ARTICLE

Mortality in trauma patients admitted during, before, and after national academic emergency medicine and trauma surgery meeting dates in Japan

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Abstract

Annually, many physicians attend national academic meetings. While participating in these meetings can have a positive impact on daily medical practice, attendance may result in reduced medical staffing during the meeting dates. We sought to examine whether there were differences in mortality after trauma among patients admitted to the hospital during, before, and after meeting dates. Using the Japan Trauma Data Bank, we analyzed in-hospital mortality in patients with traumatic injury admitted to the hospital from 2004 to 2015 during the dates of two national academic meetings—the Japanese Association for Acute Medicine (JAAM) and the Japanese Association for the Surgery of Trauma (JAST). We compared the data with that of patients admitted with trauma during identical weekdays in the weeks before and after the meetings, respectively. We used multiple logistic regression analysis to compare outcomes among the three groups. A total of 7,491 patients were included in our analyses, with 2,481, 2,492, and 2,518 patients in the during, before, and after meeting dates groups, respectively; their mortality rates were 7.3%, 8.0%, and 8.5%, respectively. After adjusting for covariates, no significant differences in in-hospital mortality were found among the three groups (adjusted odds ratio [95% CI] of the before meeting dates and after meeting dates groups; 1.18 [0.89–1.56] and 1.23 [0.93–1.63], respectively, with the during meeting dates group as the reference category). No significant differences in in-hospital mortality were found among trauma patients admitted during, before, and after the JAAM and JAST meeting dates.

Introduction

Appropriate medical staffing is essential to provide optimal trauma care [1]. Weekend or off-hours admission has been shown to be associated with worse outcomes in patients with acute myocardial infarction (AMI), stroke, pulmonary embolism, or those who required emergency

general surgery and were admitted to the intensive care unit [2–6]. This so-called “weekend effect” could possibly be explained by reduced medical staffing and resources [7, 8].

This “national meeting effect” has been examined in recent years [9–12]. Each year, many physicians attend national academic meetings and conferences to present their work, gain new knowledge, and network. Although hospitals aim to consistently deliver high quality patient care through efficient allocation of staff physicians, medical staffing during national meetings dates may be lower than that during non-meeting dates. The “national meeting effect” in Japan has been investigated; no significant differences were observed in outcomes among patients hospitalized with AMI or cardiac arrest between meeting dates and non-meeting dates [9, 10]. Interestingly, lower 30-day mortality was found among high-risk patients with AMI, cardiac arrest, and heart failure in teaching hospitals in the United States during national cardiology meeting dates [11, 12].

Although a “weekend effect” in terms of mortality has not been detected [13–15], longer emergency department stay and increased risk for missed injuries have been demonstrated for trauma patients admitted during off-hours in a community hospital setting [16]. To our knowledge, the “national meeting effect” among trauma patients has never been well elucidated. We hypothesized that hospital mortality would be higher during the meeting dates of national scientific emergency medicine and trauma surgery professional organizations than non-meeting dates and hospital mortality would be lower after the meeting dates than before the meeting dates because of reduced staffing and the positive impact of the academic meeting on high physician performance. Our study’s aim was to compare hospital mortality after trauma among patients admitted during, before, and after national meeting dates.

Materials and methods

Study design and data sources, and setting

The Okayama University Hospital ethical committee approved the study (ID 1805–020). Since patient data was extracted anonymously, the requirement for informed consent was waived. This study was designed as a nationwide retrospective cohort study. We used data from the Japan Trauma Data Bank (JTDB), which was established in 2003 with the Committee for Clinical Care Evaluation of the Japanese Association for Acute Medicine and the Trauma Surgery Committee of the Japanese Association for the Surgery of Trauma (JAST). Patients with Abbreviated Injury Scale (AIS) scores of 3 or above are recorded in the database from 264 Japanese hospitals participating in trauma research and care [17]. The registry database contains patient demographics, mechanism of injury, vital signs at the scene and on arrival, admission date, AIS scores, Injury Severity Score (ISS), treatments, and survival status at discharge from hospitals.

Each hospital provides trauma care according to local and regional trauma service systems. In our country, not all the institutions have 24/7 in-house attending trauma surgeons, nor interventional radiologist. Hence, the on-call attending trauma surgeons or interventional radiologist aim to respond immediately to the trauma call to deliver consistent quality of trauma care in these hospitals on nights and weekends [16].

Study sample

We obtained annual national meeting dates of two academic organizations—the Japanese Association for Acute Medicine (JAAM) and JAST—from 2004 to 2015. The during meeting dates group included patients admitted after traumatic injury during the dates of these meetings. The before and after meeting dates groups were defined as patients admitted with trauma during the same weekdays in the weeks before and after the meetings, respectively [9, 10]. The

JAAM and JAST meetings are each usually held for two or three consecutive days. For example, the 2015 JAAM meeting was held from Wednesday, October 21 through Friday, October 23; the before and after meeting dates groups included patients admitted Wednesday through Friday in the weeks before and after the meeting, respectively. In this study, patients who were 16 years of age or older admitted with traumatic injury from 2004 to 2015 were enrolled. Patients in cardiac arrest at the scene or on arrival and those without age, hospital arrival date, and in-hospital mortality data were excluded.

Outcome measures

Our primary outcome was post-trauma in-hospital mortality from all causes among patients hospitalized during, before, and after national meeting dates. The secondary outcome measure was hospital length of stay.

Statistical analysis

Comparisons among the three groups were made using the chi-square test for categorical variables and analysis of variance for continuous variables. We used multiple logistic regression analysis to compare the primary outcomes between the three groups, with the during meeting dates group as the reference category. Adjusted odds ratios (ORs) and their 95% confidence intervals (CIs) were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥ 65); gender; mechanism of injury (blunt or others); transfer from outside hospitals; ISS (≤ 8 vs. 9–15 vs. ≥ 16), presence or absence of shock, which was defined as systolic blood pressure of < 90 mmHg vs. 90 mmHg \leq on arrival; Glasgow Coma Scale score (≤ 8 vs. 9–15); presence or absence of emergency surgical or hemostatic intervention (craniotomy, thoracotomy, laparotomy, or angioembolization); and type of institution (high vs. low volume centers). We also applied multiple linear regression analysis adjusting for the same set of clinical variables to examine the national meeting effects on hospital length of stay. Because outcomes would be better at the high-volume centers, an additional analysis was conducted by dividing the patients into two groups; high volume centers ($\geq 1,200$ cases with ISS ≥ 9 registered for 12 years) and low volume centers ($< 1,200$ cases with ISS ≥ 9 registered for 12 years) [18]. A subgroup analysis was also conducted, stratifying patients with or without shock, polytrauma (AIS ≥ 3 for at least two different body regions) [19], and the type of national meeting. Sensitivity analysis was conducted using alternative definitions of the before and after meeting dates groups; two, three, and four weeks before and two, three, and four weeks after meeting dates, respectively, instead of one week. A two-tailed P value of < 0.05 was considered statistically significant. All analyses were performed using IBM SPSS Statistics 25 (IBM SPSS, Chicago, IL, U.S.A.).

Results

Patient characteristics

A total of 236,698 trauma patients were registered in the JTDB during the study period. Of those, 182,877 adult trauma patients were assessed for eligibility. After 175,386 patients were excluded due to not being admitted on eligible days, 7,491 subjects were included in our analyses, with 2,481 patients in the during meeting dates group, 2,492 patients in the before meeting dates group, and 2,518 patients in the after meeting dates group (Fig 1). Among the three groups of patients, basic characteristics including severity of trauma and life-saving surgical procedures were similar except for the age category (Table 1).

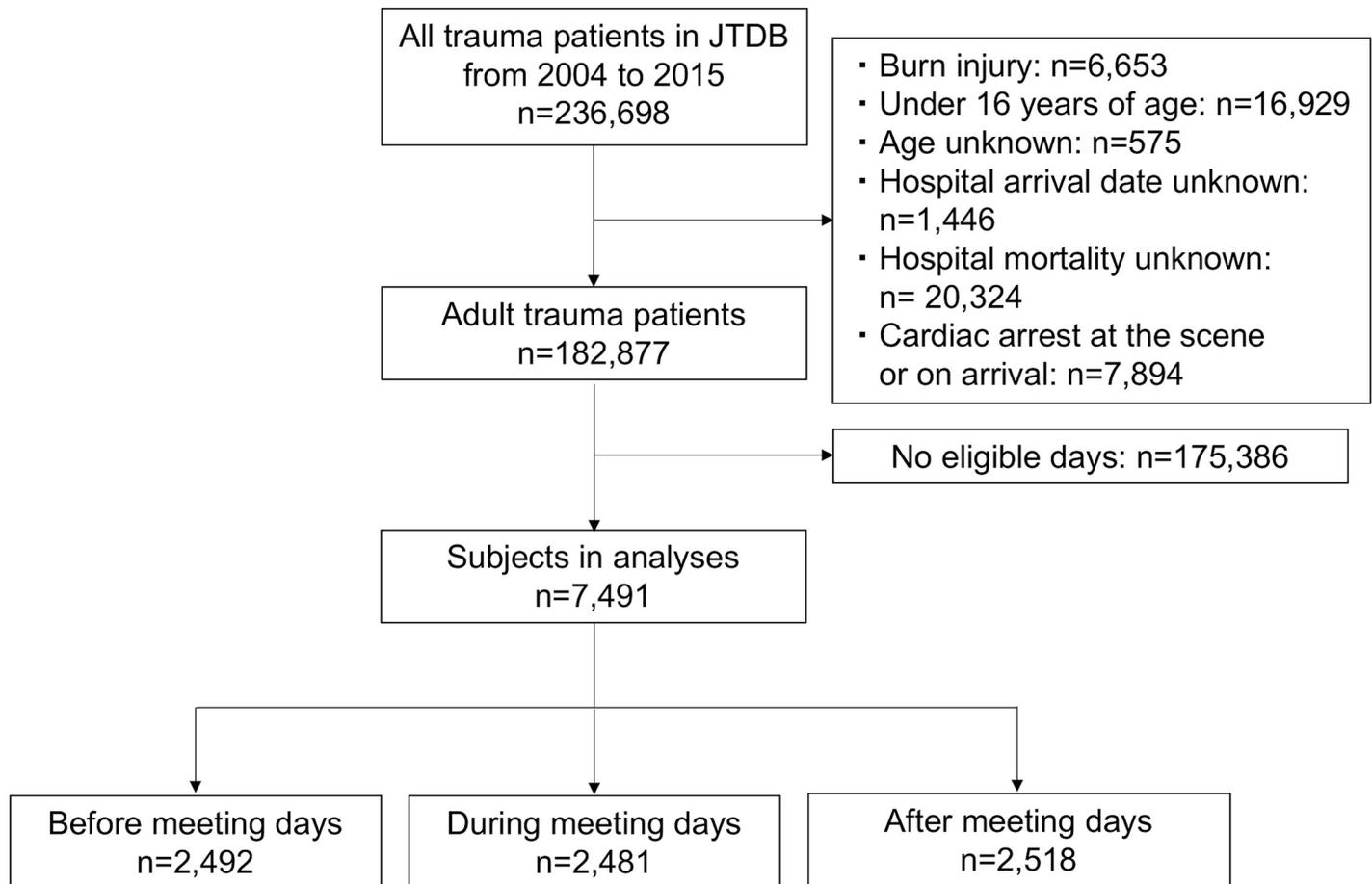


Fig 1. Flow diagram of the study population. JTDB, Japan Trauma Data Bank.

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Comparison of mortality between the three groups

No significant differences in in-hospital mortality were observed during, before, and after meeting dates (7.3% vs. 8.0% vs. 8.5%, respectively; $P = 0.306$; unadjusted OR [95% CI] of the before and after meeting dates groups; 1.11 [0.90–1.37], 1.17 [0.96–1.44], respectively, with the during meeting dates group as the reference category; Table 2). Even after adjusting for covariates, no significant differences in in-hospital mortality were found among the three groups (adjusted OR [95% CI] of the before and after meeting dates groups; 1.18 [0.89–1.56], 1.23 [0.93–1.63], respectively, with the during meeting dates group as the reference category; Table 2).

Subgroup analysis

Although high volume centers were associated with better outcomes (6.8% overall in-hospital mortality of high volume centers vs. 9.5% for the low volume centers; $P < 0.001$), in-hospital mortality did not differ among the three groups according to center volume (Table 3). Additional analyses were conducted by stratifying patients with or without shock, and polytrauma; no differences in in-hospital mortality were found among the three groups (Table 3). Also, no significant differences according to type of national meeting were found (Table 3). The same results were obtained when considering alternative definitions of the before and after meeting dates groups (Table 4).

Table 1. Characteristics of trauma injury patients admitted during, before, and after national meeting dates.

| | Before meeting dates group n = 2,492 | During meeting dates group n = 2,481 | After meeting dates group n = 2,518 | P-Value |
|--------------------------------------------|-----------------------------------------|-----------------------------------------|----------------------------------------|---------|
| Age (years), median (IQR) | 63 (41, 77) | 63 (41, 77) | 65 (43, 78) | 0.131 |
| 16–39 (years), n (%) | 541 (21.7) | 604 (24.3) | 566 (22.5) | 0.015 |
| 40–64 (years), n (%) | 750 (30.1) | 708 (28.5) | 675 (26.8) | |
| ≥65 (years), n (%) | 1,201 (48.2) | 1,169 (47.1) | 1,277 (50.7) | |
| Male, n (%) | 1,557 (62.5) | 1,513 (61.1) | 1,607 (63.9) | 0.126 |
| Blunt mechanism, n (%) | 2,347 (94.2) | 2,348 (94.6) | 2,398 (95.2) | 0.249 |
| Transfers from an outside hospital, n (%) | 345 (13.8) | 367 (14.8) | 345 (13.7) | 0.485 |
| SBP (mmHg), median (IQR) | 136 (116, 158) | 136 (116, 156) | 137 (117, 158) | 0.624 |
| <90mmHg, n (%) | 180 (7.2) | 180 (7.3) | 170 (6.8) | 0.731 |
| Missing, n (%) | 60 (2.4) | 54 (2.2) | 55 (2.2) | |
| Glasgow Coma Scale, median (IQR) | 15 (14, 15) | 15 (14, 15) | 15 (14, 15) | 0.532 |
| ≤8, n (%) | 245 (9.8) | 242 (9.8) | 256 (10.1) | 0.874 |
| Missing, n (%) | 236 (9.5) | 193 (7.8) | 211 (8.4) | |
| Surgical or hemostatic intervention, n (%) | 266 (10.7) | 237 (9.6) | 237 (9.4) | 0.261 |
| Craniotomy, n (%) | 101 (4.0) | 83 (3.4) | 82 (3.3) | 0.249 |
| Thoracotomy, n (%) | 32 (1.3) | 31 (1.3) | 20 (0.8) | 0.181 |
| Laparotomy, n (%) | 76 (3.1) | 72 (2.9) | 66 (2.6) | 0.652 |
| Angioembolization, n (%) | 73 (2.9) | 65 (2.6) | 78 (3.1) | 0.592 |
| Injury Severity Score, median (IQR) | 10 (9, 20) | 10 (9, 20) | 10 (9, 21) | 0.590 |
| Injury Severity Score, mean (SD) | 15 (11) | 15 (11) | 15 (11) | |
| ≤8, n (%) | 407 (16.3) | 399 (16.1) | 398 (15.8) | 0.904 |
| 9–15, n (%) | 993 (39.8) | 980 (39.5) | 1,016 (40.3) | |
| ≥16, n (%) | 1,009 (40.5) | 1,042 (42.0) | 1,034 (41.1) | |
| Missing, n (%) | 83 (3.3) | 60 (2.4) | 70 (2.8) | |
| Polytrauma, n (%) | 400 (16.1) | 446 (18.0) | 451 (17.9) | 0.125 |
| High volume center, n (%) | 1,437 (57.7) | 1,486 (59.9) | 1,450 (57.6) | 0.172 |
| Low volume center, n (%) | 1,055 (42.3) | 995 (40.1) | 1,068 (42.4) | |
| Hospital LOS (days), median (IQR) | 16 (5, 32) | 15 (5, 31) | 16 (5, 32) | 0.479 |
| Missing, n (%) | 28 (1.1) | 36 (1.5) | 22 (0.9) | |
| In-hospital mortality, n (%) | 200 (8.0) | 181 (7.3) | 213 (8.5) | 0.306 |

IQR: interquartile range; SBP: systolic blood pressure; SD: standard deviation; LOS: length of stay.

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Table 2. In-hospital mortality after trauma among patients hospitalized during, before, and after national meeting dates.

| | Before meeting dates group | During meeting dates group | After meeting dates group | P-value |
|--------------------------------|----------------------------|----------------------------|---------------------------|---------|
| Overall | | | | |
| In-hospital mortality, % (n/N) | 8.0 (200/2,492) | 7.3 (181/2,481) | 8.5 (213/2,518) | 0.306 |
| Crude OR (95% CIs) | 1.11 (0.90–1.37) | Reference | 1.17 (0.96–1.44) | |
| Adjusted OR (95% CIs) | 1.18 (0.89–1.56) | Reference | 1.23 (0.93–1.63) | |

Adjusted OR and their 95% CIs were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥65), gender, mechanism of injury (blunt or others), transfer from an outside hospital, ISS (≤8 vs. 9–15 vs. ≥16), presence or absence of shock, Glasgow Coma Scale score (≤8 vs. 9–15), presence or absence emergency surgical or hemostatic intervention, and type of institution (high vs. low volume centers).

OR: odds ratio; CI: confidence interval; ISS: Injury Severity Score

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Secondary outcome

Consistent with a non-significant difference in the hospital length of stay among the three groups in a descriptive analysis (Table 1), multiple linear regression analysis did not show any

Table 3. In-hospital mortality among the three groups with stratification for high vs. low volume centers, presence or absence of shock, polytrauma, and type of national meeting.

| | Before meeting dates group | During meeting dates group | After meeting dates group | P-value |
|--------------------------------------------|----------------------------|----------------------------|---------------------------|---------|
| High volume centers | | | | |
| In-hospital mortality, % (n/N) | 7.2 (103/1,437) | 6.0 (89/1,486) | 7.3 (106/1,450) | 0.129 |
| Crude OR (95% CI) | 1.21 (0.90–1.62) | Reference | 1.24 (0.93–1.66) | |
| Adjusted OR (95% CI) ^a | 1.12 (0.77–1.64) | Reference | 1.30 (0.89–1.89) | |
| Low volume centers | | | | |
| In-hospital mortality, % (n/N) | 9.2 (97/1,055) | 9.2 (92/995) | 10.0 (107/1,068) | 0.968 |
| Crude OR (95% CI) | 0.99 (0.74–1.34) | Reference | 1.09 (0.82–1.47) | |
| Adjusted OR (95% CI) ^a | 1.25 (0.82–1.90) | Reference | 1.17 (0.77–1.78) | |
| Systolic blood pressure ≤ 90mmHg | | | | |
| In-hospital mortality, % (n/N) | 29.4 (53/180) | 35.6 (64/180) | 34.7 (59/170) | 0.413 |
| Crude OR (95% CI) | 0.76 (0.49–1.18) | Reference | 0.96 (0.62–1.50) | |
| Adjusted OR (95% CI) ^b | 1.09 (0.62–1.90) | Reference | 1.14 (0.66–1.97) | |
| Systolic blood pressure > 90mmHg | | | | |
| In-hospital mortality, % (n/N) | 6.0 (134/2,252) | 4.5 (101/2,247) | 6.0 (138/2,293) | 0.040 |
| Crude OR (95% CI) | 1.34 (1.03–1.75) | Reference | 1.36 (1.05–1.77) | |
| Adjusted OR (95% CI) ^b | 1.28 (0.93–1.74) | Reference | 1.29 (0.95–1.77) | |
| Polytrauma | | | | |
| In-hospital mortality, % (n/N) | 21.5 (86/400) | 20.6 (92/446) | 22.2 (100/451) | 0.852 |
| Crude OR (95% CI) | 1.05 (0.76–1.47) | Reference | 1.10 (0.80–1.51) | |
| Adjusted OR (95% CI) ^c | 1.32 (0.87–1.98) | Reference | 1.23 (0.83–1.84) | |
| JAAM | | | | |
| In-hospital mortality, % (n/N) | 8.3 (130/1,560) | 7.4 (112/1,513) | 9.1 (147/1,612) | 0.221 |
| Crude OR (95% CI) | 1.14 (0.87–1.48) | Reference | 1.26 (0.97–1.62) | |
| Adjusted OR (95% CI) ^d | 1.28 (0.90–1.80) | Reference | 1.28 (0.91–1.80) | |
| JAST | | | | |
| In-hospital mortality, % (n/N) | 7.5 (70/932) | 7.1 (69/968) | 7.3 (66/906) | 0.950 |
| Crude OR (95% CI) | 1.06 (0.75–1.49) | Reference | 1.02 (0.72–1.45) | |
| Adjusted OR (95% CI) ^d | 1.21 (0.78–1.88) | Reference | 1.18 (0.75–1.87) | |

^aAdjusted OR and their 95% CIs were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥65), gender, mechanism of injury (blunt or others), transfer from an outside hospital, ISS (≤8 vs. 9–15 vs. ≥16), presence or absence of shock, Glasgow Coma Scale score (≤8 vs. 9–15), and presence or absence of emergency surgical or hemostatic intervention.

^bAdjusted OR and their 95% CIs were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥65), gender, mechanism of injury (blunt or others), transfer from an outside hospital, ISS (≤8 vs. 9–15 vs. ≥16), Glasgow Coma Scale score (≤8 vs. 9–15), presence or absence of emergency surgical or hemostatic intervention, and type of institution (high vs. low volume centers).

^cAdjusted OR and their 95% CIs were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥65), gender, mechanism of injury (blunt or others), transfer from an outside hospital, presence or absence of shock, Glasgow Coma Scale score (≤8 vs. 9–15), presence or absence of emergency surgical or hemostatic intervention, and type of institution (high vs. low volume centers).

^dAdjusted OR and their 95% CIs were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥65), gender, mechanism of injury (blunt or others), transfer from an outside hospital, ISS (≤8 vs. 9–15 vs. ≥16), presence or absence of shock, Glasgow Coma Scale score (≤8 vs. 9–15), presence or absence of emergency surgical or hemostatic intervention, and type of institution (high vs. low volume centers).

OR: odds ratio; CIs: confidence intervals; ISS: Injury Severity Score; JAAM, Japanese Association for Acute Medicine; JAST, Japanese Association for the Surgery of Trauma.

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Table 4. In-hospital mortality among the three groups with alternative definitions of the before and after meeting dates groups.

| | Before meeting dates group | During meeting dates group | After meeting dates group | P-value |
|--------------------------------|----------------------------|----------------------------|---------------------------|---------|
| $\pm 2^a$ | | | | |
| In-hospital mortality, % (n/N) | 7.1 (169/2,397) | 7.3 (181/2,481) | 7.5 (190/2,544) | 0.851 |
| Adjusted OR (95% CIs) | 1.07 (0.80–1.42) | Reference | 1.27 (0.96–1.68) | |
| $\pm 3^b$ | | | | |
| In-hospital mortality, % (n/N) | 7.6 (193/2,546) | 7.3 (181/2,481) | 8.2 (207/2,517) | 0.450 |
| Adjusted OR (95% CIs) | 1.22 (0.92–1.61) | Reference | 1.21 (0.92–1.61) | |
| $\pm 4^c$ | | | | |
| In-hospital mortality, % (n/N) | 7.7 (196/2,538) | 7.3 (181/2,481) | 6.6 (176/2,650) | 0.316 |
| Adjusted OR (95% CIs) | 1.23 (0.92–1.62) | Reference | 0.94 (0.71–1.26) | |

^aTwo weeks before and after meeting dates as the before meeting dates group and after meeting dates group.

^bThree weeks before and after meeting dates as the before meeting dates group and after meeting dates group.

^cFour weeks before and after meeting dates as the before meeting dates group and after meeting dates group.

Adjusted OR and their 95% CIs were obtained after adjusting for age (16–39 vs. 40–64 vs. ≥ 65), gender, mechanism of injury (blunt or others), transfer from an outside hospital, ISS (≤ 8 vs. 9–15 vs. ≥ 16), presence or absence of shock, Glasgow Coma Scale score (≤ 8 vs. 9–15), presence or absence of emergency surgical or hemostatic intervention, and type of institution (high vs. low volume centers).

OR: odds ratio; CIs: confidence intervals; ISS: Injury Severity Score.

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significant difference in the hospital length of stay between the three groups. Indeed, compared with the reference category (during meeting dates group), the beta coefficients were 16.61 (95% CI: -6.74 to 39.98) for before meeting dates group and -1.72 (95% CI: -25.00 to 21.56) for after meeting dates group, respectively.

Discussion

In this study, we investigated whether there was a difference in mortality among patients admitted due to traumatic injuries during, before, and after dates of national academic acute medicine and trauma meetings. Contrary to our hypothesis, we found no significant differences in in-hospital mortality among the three groups, even after adjusting for measurable confounders.

To our knowledge, “national meeting effects” were first investigated in the United States, focusing on national cardiology meetings; lower staffing and differences in composition of physicians during the meeting dates were found to possibly affect treatment utilization and outcomes [11]. In this study, no significant differences in mortality of AMI patients were found between those in the hospital during meeting and non-meeting dates; however, high-risk patients with AMI, cardiac arrest, and heart failure admitted to teaching hospitals during meeting dates were found to have lower mortality than those admitted during non-meeting dates [11]. The present study is the first to examine the “national meeting effect” regarding mortality in trauma patients.

Previous studies have not detected the “weekend effect”; admission on nights or weekends for trauma patients was not associated with increased mortality [13–16, 20, 21] or even better outcomes, which was explained by the possibility that the patients could have been immediately accessed to the operation room or resources that might otherwise be occupied during normal working hours [22]. Generally, a plausible explanation for the “weekend effect” includes several factors such as reduced medical staffing, decreased access to some tests and procedures, and the influence of variations in case mix [20, 21, 23]. For trauma patients, hospitals are explicitly required to be appropriately staffed and to provide optimal care, regardless of

when injured patients are admitted [22, 24]. However, trauma patients presenting off-hours were more likely to have missed injuries, [16, 25], in particular thoracic spine or abdominal injuries [25]. Specifically, Schwartz DA, et al. showed that off-hour presentation of pelvic fracture patients with hemorrhagic shock caused a delay in door-to-angiobolization time, resulting in increased mortality [24].

In theory, variations of patient characteristics would not be influenced by admission dates occurring during, before, and after national meetings. While relatively less experienced emergency physicians or trauma surgeons providing leadership would be expected during academic national meeting dates, strong leadership, teamwork, and technical skills are essential components for team performance and patient care in initial trauma management [26, 27]. Our results showed no significant differences in mortality among those admitted during, before, and after national meeting dates; these findings may be explained by the possibility that every hospital strives to consistently deliver high quality care. In those hospitals where attending trauma surgeons are unavailable 24/7, on-call attending staffs aim to respond immediately to the trauma call to deliver consistent quality of trauma care during off-hour [16]. Similarly, each hospital is assumed to properly allocate and cover attending staffs during the national meeting dates. A previous study regarding national meeting effect, in which results demonstrated better outcomes among high-risk patients with heart failure or cardiac arrest, were possibly explained by declines in intensity of care or volume of less urgent cardiovascular hospitalization during meetings dates [11]. However, these explanations cannot be applied to our results because we found no evidence that total trauma patient volume or life-saving interventions for trauma patients declined during the national meeting dates. In this study, detailed information was unavailable on rostered attending staffs either in-house or on-call during, before, and after national meeting dates. A working pattern completely depended on each institution's policy and physicians who remain to treat trauma patients would concentrate more attention on the patients, by which clinical outcomes were not influenced by the national meeting.

Although no clear evidence exists that national academic meetings directly improve medical staff performance and positively impact patient outcomes, hospital participation in a trauma quality improvement program has been demonstrated to be associated with better patient outcomes [28]. Hence, we compared hospital mortality among patients admitted during, before, and after national meeting dates to investigate the "post-national meeting effect," assuming that participation in national academic meetings has a positive effect on clinical performance and patient outcomes. Despite our hypothesis, our results detected no "post-national meeting effect."

Our study has several limitations. First, we could not identify differences in medical staffing among the three groups, such as composition of emergency physicians vs. trauma surgeons who treated the patients; therefore, we are unable to explain why no significant differences in mortality were found. Second, the influence of national meeting geographical regions and locations was not accounted for, which may have affected our results. Third, as Jena AB, et al. showed that high-risk AMI patients admitted to teaching hospitals during national meeting dates received less percutaneous coronary intervention, hospital type should have been taken into consideration [11]. We performed additional analyses according to hospital volume, in which the same results were obtained. Fourth, possible confounders, including comorbidities, were unavailable in this study. Fifth, although major tertiary hospitals providing high-quality trauma care participate in the JTDB, a degree of random error and selection bias may have occurred, as this was not a comprehensive study [17]. Finally, since we focused on national academic meetings only in Japan, our results could not be applied to other countries, considering differences in the settings and geography of Japanese healthcare systems.

Conclusions

We observed no significant differences in in-hospital mortality after trauma among patients admitted during, before, and after national acute medicine and trauma meeting dates. As hospitals are assumed to be struggling to consistently provide optimal care for trauma patients, participating in these meetings is acceptable for sharing and generating new knowledge. Further population-based studies are required to validate our results.

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